

Meta-Regression

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Meta-Regression

- In case of substantial heterogeneity between the studies, possible causes of the heterogeneity should be explored.
- In the context of meta-analysis this can be done by either covariates on the study level that could explain the differences between the studies or by covariates on the subject level.
- However, the latter approach is only possible when individual data are available.
- Since often only information on the study level is available, explaining and investigating heterogeneity by covariates on the study level has drawn much attention in applied sciences.

Meta-Regression

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Meta-Regression

- Since the number of studies in a meta-analysis is usually quite small, there is a great danger of overfitting.
- So, there is only room for a few explanatory variables in a meta-regression, whereas a lot of characteristics of the studies may be identified as potential causes of heterogeneity.
- Investigations of differences between the studies and their results are observational associations and are subject to biases (such as aggregation bias) and confounding (resulting from correlation between study characteristics).
- Consequently, there is a clear danger of misleading conclusions if P -values from multiple meta-regression analyses are interpreted naïvely.

Meta-Regression

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Example

Data on 13 trials on the prevention of tuberculosis using BCG vaccination

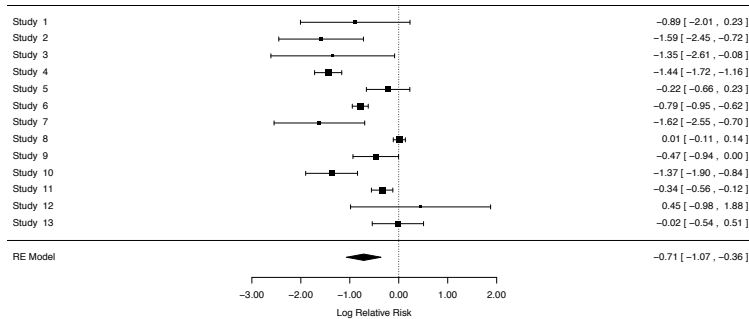
Trial	Vaccinated		Not vaccinated		Latitude
	Disease	No disease	Disease	No Disease	
1	4	119	11	128	44
2	6	300	29	274	55
3	3	228	11	209	42
4	62	13536	248	12619	52
5	33	5036	47	5761	13
6	180	1361	372	1079	44
7	8	2537	10	619	19
8	505	87886	499	87892	13
9	29	7470	45	7232	27*
10	17	1699	65	1600	42
11	186	50448	414	27197	18
12	5	2493	3	2338	33
13	27	16886	29	17825	33

Further covariates available: Year of publication, type of allocation (alternate, random, systematic)

Meta-Regression

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Example



Meta-Regression (one covariate)

Let us consider k independent trials (experiments) and each trial provides an estimate, say $\hat{\theta}_i$, $i = 1, \dots, k$, of a parameter of interest, say θ , and an estimate of the variance of $\hat{\theta}_i$, say $\hat{\sigma}_i^2$, $i = 1, \dots, k$. Moreover, one covariate on study-level, say x_i , is known for each trial.

Normal-normal hierarchical model:

$$\hat{\theta}_i \sim N(\theta_i, \hat{\sigma}_i^2)$$

and

$$\theta_i \sim N(\theta_{MA}, \tau_{MA}^2)$$

OR

$$\theta_i \sim N(\theta_{MR} + \beta x_i, \tau_{MR}^2)$$

Meta-Regression (one covariate)

Random effects meta-analysis model

$$\hat{\theta}_i \sim N(\theta_{MA}, \tau_{MA}^2 + \hat{\sigma}_i^2).$$

- θ_{MA} – mean effect size
- τ_{MA}^2 – between-study variability (heterogeneity parameter)

Random effects meta-regression

$$\hat{\theta}_i \sim N(\theta_{MR} + \beta x_i, \tau_{MR}^2 + \hat{\sigma}_i^2).$$

- θ_{MR} – effect size given that the covariate is zero
- τ_{MR}^2 – residual heterogeneity

Meta-Regression (one covariate)

Random effects meta-regression

$$\hat{\theta}_i \sim N(\theta_{MR} + \beta x_i, \tau_{MR}^2 + \hat{\sigma}_i^2).$$

Objectives:

- Fixed effects or random effects meta-regression?
Test of $H_0 : \tau_{MR}^2 = 0$!
- Estimate and confidence interval for τ_{MR}^2
- Estimates and confidence intervals for θ_{MR} and β

Extend analysis methods from meta-analysis to meta-regression:

- (conditional) restricted maximum likelihood estimation
- method of moments estimation

R Package metafor

- "Classical" meta-analysis and meta-regression (weighted least squares method)
- Knapp-Hartung (2003) approach for meta-analysis and meta-regression
- Q -profiling confidence interval for τ^2
- A lot of estimators for the heterogeneity parameter τ^2
- . . .

Note: The *metareg* function in R package *meta* can also be used (wrapper function that calls *rma.uni* function from R package *metafor*).

Inference on the Fixed Effects

- Hence, unbiased and non-negative estimators of the variances of $\tilde{\theta}$ and $\tilde{\beta}$ are given by

$$Q_2(\tilde{\theta}) = \frac{1}{k-2} \sum_{i=1}^k g_i (Y_i - \tilde{\theta} - \tilde{\beta} x_i)^2$$

with $g_i = w_i / [\sum w_j - (\sum w_j x_j)^2 / \sum w_j x_j^2]$, $i = 1, \dots, k$, and

$$Q_2(\tilde{\beta}) = \frac{1}{k-2} \sum_{i=1}^k h_i (Y_i - \tilde{\theta} - \tilde{\beta} x_i)^2$$

with $h_i = w_i / [\sum w_j x_j^2 - (\sum w_j x_j)^2 / \sum w_j]$, $i = 1, \dots, k$.

Inference on the Fixed Effects

- Let $\tilde{\theta}$ and $\tilde{\beta}$ be the weighted least-squares estimators with known variances.
- Knapp and Hartung (2003) considered the quadratic form

$$Q_2 = \frac{1}{k-2} \sum_{i=1}^k w_i (Y_i - \tilde{\theta} - \tilde{\beta} x_i)^2, \quad k > 2.$$

that is, a mean sum of the weighted least-squares residuals.

- Under normality of Y_i , the quadratic form Q_2 is stochastically independent of the $\tilde{\theta}$ and $\tilde{\beta}$, and $(k-2) Q_2$ is χ^2 -distributed with $k-2$ degrees of freedom.

Inference on the Fixed Effects

- Replacing the unknown variance components in $Q_2(\tilde{\theta})$ and $Q_2(\tilde{\beta})$ by appropriate estimates, Knapp and Hartung (2003) proposed the following approximate $(1-\alpha)$ -confidence intervals on θ and β :

$$\hat{\theta} \pm \sqrt{\hat{Q}_2(\hat{\theta})} t_{k-2; \alpha/2}$$

and

$$\hat{\beta} \pm \sqrt{\hat{Q}_2(\hat{\beta})} t_{k-2; \alpha/2}.$$

Example

Results for slope with covariate latitude

Method $\hat{\tau}^2$	Estimate	95% CI (classical)	95% CI (KH)
Hunter-Schmidt	-0.0296	[-0.0398, -0.0193]	[-0.0447, -0.0144]
Hedges	-0.0282	[-0.0489, -0.0075]	[-0.0493, -0.0071]
DerSimonian-Laird	-0.0292	[-0.0424, -0.0160]	[-0.0467, -0.0118]
Sidik-Jonkman	-0.0281	[-0.0497, -0.0065]	[-0.0495, -0.0067]
ML	-0.0295	[-0.0403, -0.0188]	[-0.0452, -0.0139]
REML	-0.0291	[-0.0432, -0.0150]	[-0.0472, -0.0111]
Paule-Mandel	-0.0286	[-0.0463, -0.0108]	[-0.0485, -0.0086]

Explaining Heterogeneity

Method $\hat{\tau}^2$	MA	MR	Reduction (in %)
Hunter-Schmidt	0.2284	0.0291	87.26
Hedges	0.3285	0.2090	36.38
DerSimonian-Laird	0.3087	0.0633	79.50
Sidik-Jonkman	0.3455	0.2318	32.90
ML	0.2800	0.0344	87.73
REML	0.3132	0.0764	75.62
Paule-Mandel	0.3180	0.1421	55.31

Categorical Covariate

```
### load package
load(metafor)
### load BCG vaccine data
data(dat.bcg)
### calculate log relative risks and corresponding sampling variances
dat <- escalc(measure="RR", ai=tpos, bi=tneg, ci=cpos,
              di=cneg, data=dat.bcg)
### using a model formula to specify the same model
rma(yi, vi, mods = ~ factor(alloc), data=dat, method="REML", btt=c(2,3))
```

Categorical Covariate

Mixed-Effects Model (k = 13; tau² estimator: REML)

```
tau^2 (estimated amount of residual heterogeneity):      0.3615
tau (square root of estimated tau^2 value):              0.6013
I^2 (residual heterogeneity / unaccounted variability): 88.77%
H^2 (unaccounted variability / sampling variability):     8.91
R^2 (amount of heterogeneity accounted for):              0.00%
```

```
Test for Residual Heterogeneity:
QE(df = 10) = 132.3676, p-val < .0001
```

```
Test of Moderators (coefficient(s) 2,3):
QM(df = 2) = 1.7675, p-val = 0.4132
```

Categorical Covariate

Model Results:

	estimate	se	zval	pval	ci.lb	ci.ub
intrcpt	-0.5180	0.4412	-1.1740	0.2404	-1.3827	0.3468
random	-0.4478	0.5158	-0.8682	0.3853	-1.4588	0.5632
systematic	0.0890	0.5600	0.1590	0.8737	-1.0086	1.1867

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Results for type of allocation

Further Applications

- Methods can be easily extended to more than one covariate.
- Meta-regression is primarily used for explaining heterogeneity between study results.
- Meta-regression technique can be also used for other applications of combining results; e.g. combining results from controlled and uncontrolled studies in meta-regression model where the covariate indicates whether the result comes from a controlled or from an uncontrolled study